

Some Applications of Mixed Volumes in Data Science

Eliza O'Reilly

Caltech

Application #1: Prediction with Random Tessellation Forests

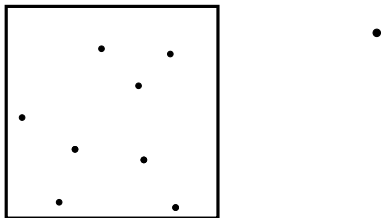
Joint work with Ngoc Mai Tran (UT Austin)

Goal: Given example input-output pairs $f(x_i; y_i)_{i=1}^n \quad \mathbb{R}^d \rightarrow \mathbb{R}$, obtain estimator \hat{f}_n to predict output from new input x : $\hat{y} = \hat{f}_n(x)$

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Randomized Decision Trees:

- Recursively split data along random feature of input
- Induce a hierarchical axis-aligned partition of input space



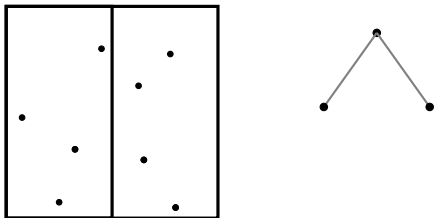
- Random tree (regression) estimator: average over cell/leaf

$$\hat{f}_n(x) = \frac{\sum_{i=1}^n y_i \mathbb{1}_{f(x_i) \text{ in same cell as } x}}{\sum_{i=1}^n \mathbb{1}_{f(x_i) \text{ in same cell as } x}}$$

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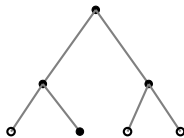
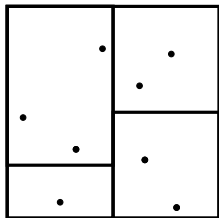
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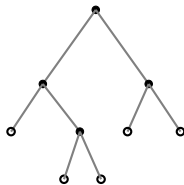
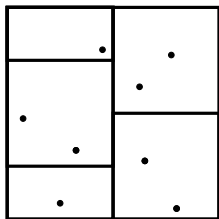
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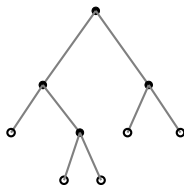
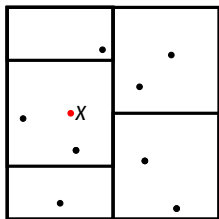
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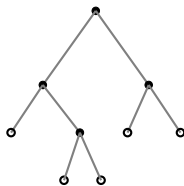
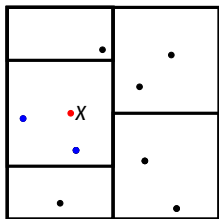
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Random Forest (RF) Estimator

- | Average of *M*i.i.d. tree estimators

¹[Ho, 1995; 1998; Amit and Geman, 1997; Breiman, 2001]

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- | State-of-the-art empirical performance for many tasks²

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- | Purely RF variants⁴: splits *independent* of data

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Mondrian process

- | Introduced by Roy and Teh in 2008
- | Stochastic process that recursively builds an axis-aligned hierarchical partition in \mathbb{R}^d

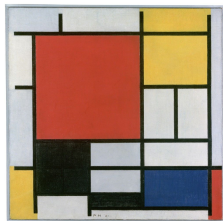
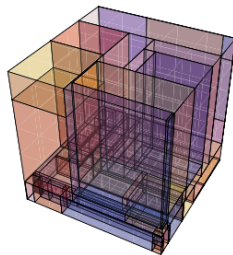
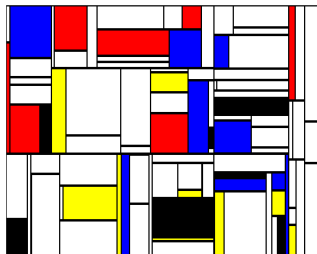


Figure: Piet Mondrian (1921).

Mondrian process construction in \mathbb{R}^d

1. Fix lifetime parameter $\tau > 0$

2. Draw

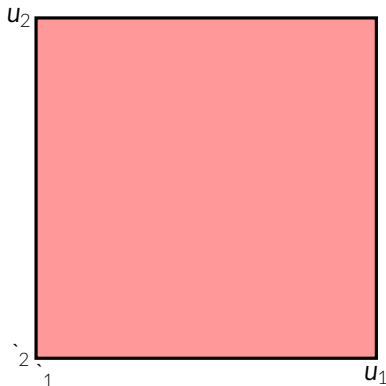
$$\text{Exp} \left(\sum_{i=1}^d u_i \right)$$

3. IF $t > \tau$ stop,

ELSE sample a split:

- *Dimension*: j with probability proportional to u_j
- *Location*: uniform on $[0, u_j]$

4. Recurse independently on each subrectangle with lifetime $\tau/4$



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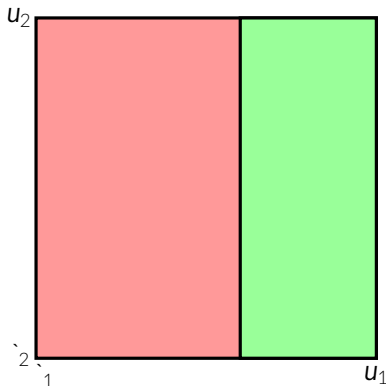
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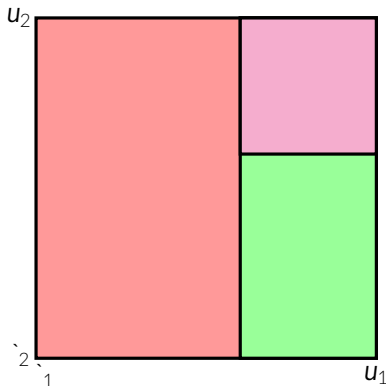
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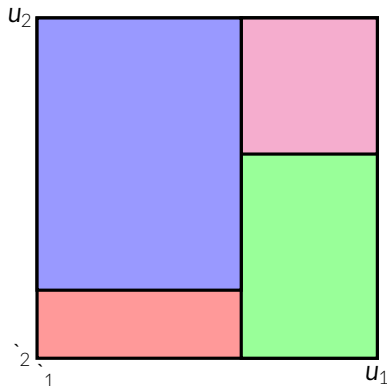
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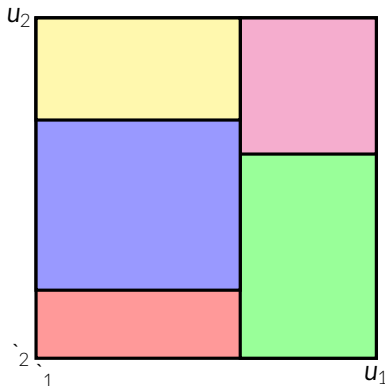
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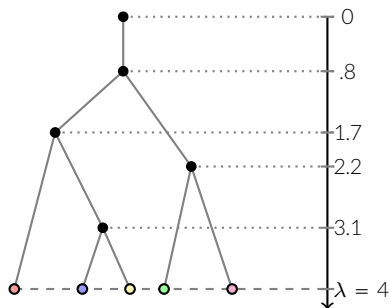
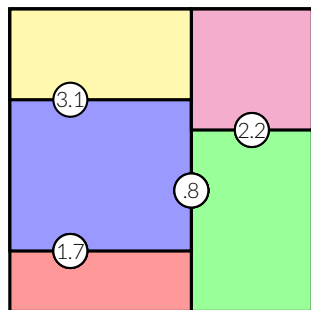
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Mondrian Random Forests

- Comparable empirical performance to RF for some tasks⁵
- Minimax rates under nonparametric assumptions in arbitrary dimension⁶



⁵[Lakshminarayanan, Roy, and Teh, 2014]

⁶[Mourtada, Gaïffas, Scornet, 2020]

Beyond axis-aligned partitions

- | Non-axis-aligned splits can capture dependencies between features
- | Non-axis-aligned RF variants show **improved empirical performance**⁷
- | Lack of theoretical analysis, computational efficiency

⁷[Breiman, 2001; Fan, Li, and Sisson, 2019; Tomita et al., 2020]

Beyond axis-aligned partitions

- | Non-axis-aligned splits can capture dependencies between features
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Question: Is there a generalization of the Mondrian process with non-axis-aligned cuts?

⁷[Breiman, 2001; Fan, Li, and Sisson, 2019; Tomita et al., 2020]

Stable Under Iteration Processes

- | **Yes!** Mondrian is special case of the **STIT process** in stochastic geometry
- | Introduced by Nagel and Weiss in 2003
- | Indexed by a **directional distribution** on S^{d-1}

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- | **Yes!** Mondrian is special case of the **STIT process** in stochastic geometry
- | Introduced by Nagel and Weiss in 2003
- | Indexed by a **directional distribution** on S^{d-1}

- | Improved empirical performance with uniform STIT over Mondrian⁸
- | General cell shapes introduce computational and **theoretical** challenges

⁸[Ge, Wang, Teh, Wang, and Elliott, 2019]

4 THEORETICAL CHALLENGE - MINIMAX STATES

THEORETICAL CHALLENGE - IN-MAXIMATES

- ASSUMPTION 1: $(X_i, Y_i)_{i=1}^N$ IID SAMPLE FROM \mathcal{P} SUCH THAT

$$g = \mathbb{E}(Y) + \epsilon;$$

WHERE \mathcal{R}^D IS A COMPACT AND CONVEX WINDOW

- \hat{g}_N - 34)4 FOREST ESTIMATOR OF MEAN PARAMETER

THEORETICAL CHALLENGE - IN-MAXIMATES

- ASSUMPTION 1: $(X_i, Y_i)_{i=1}^n$ IID SAMPLES SUCH THAT

$$Y = F(\theta) + \epsilon;$$

WHERE R^D IS A COMPACT AND CONVEX WINDOW

- $\hat{F}_{N, \lambda}$ 4 FOREST ESTIMATOR OF SIZE PARAMETER
- THE QUALITY OF THE ESTIMATION MEASURED BY THE QUADRATIC RISK

$$E[(\hat{F}_{N, \lambda}(\theta) - F(\theta))^2]$$

THEORETICAL CHALLENGE - MINIMAX RATES

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$$E[(\hat{f}_{N, \lambda}(\theta) - f(\theta))^2]$$

- THE MINIMAX RISK FOR A FUNCTION CLASS

$$\min_{\hat{f}_N} \max_{f \in \mathcal{F}} E[(\hat{f}_N(\theta) - f(\theta))^2]$$

THEORETICAL CHALLENGE - IN MAXIMATES

THEOREM 4.1 AND /

1) BY THE WEAK COMPACTNESS THEOREM OF ALFONSO DENOTING $N = (D^+)$ GIVES

$$E[(\hat{F}_{N;N}(\theta) - F(\theta))] / N = (D^+)$$

4HEORETICAL #HALLENGE -INIMAX 2ATES

4HEOREM 4RAN AND /

I)FFIS, IPSCHITZ LETTING $N = (D^+)$ GIVES

$$E[(\hat{F}_{N;N-}(\delta) \quad F(\delta))] / N = (D^+)$$

II)FFISC AND δ HAS POSITIVE AND, IPSCHITZ DENSITY LETTING AND
- N & $N = (D^+)$ GIVES

$$E[(\hat{F}_{N;N-N}(\delta) \quad F(\delta))] / N = (D^+)$$

THEORETICAL CHALLENGE - INIMAX TESTS

THEOREM 4.1 AND /

I) IF f IS μ -IPSCHITZ LETTING $N = (D^+)$ GIVES

$$E[(\hat{F}_{N;N-}(\theta) - f(\theta))] / N = (D^+)$$

II) IF f IS μ -IPSCHITZ DENSITY LETTING $N = (D^+)$ AND $N \rightarrow \infty$ GIVES

$$E[(\hat{F}_{N;N-}(\theta) - f(\theta))] / N = (D^+)$$

I) 3.4) RANDOM FORESTS INIMAX OPTIMATOR, IPSCHITZ AND FUNCTIONS

RATES FOR MULTI INDEX MODELS

3 SUPPOSE FOR S ! RANDOM D $I = ; \dots ; S$

$$F(X) = G(h_A; X; \dots; h_S; X); \quad X \in \mathbb{R}^D \quad (; 2):$$

RATES FOR MULTI INDEX MODELS

1. SUPPOSE $f: \mathbb{R}^S \rightarrow \mathbb{R}$ AND $\mathcal{X} \subset \mathbb{R}^D$ IS A SET OF DATA POINTS

$$f(x) = G(h_{A_1}(x); \dots; h_{A_S}(x)) \quad \mathcal{X} \subset \mathbb{R}^D \text{ (} D \geq S \text{)}$$

2. $\mathcal{X} := \text{SPAN}(A_1; \dots; A_S)$ \mathbb{R}^D IS THE S -DIMENSIONAL RELEVANT FEATURE SUBSPACE

2ATES FOR MULTI INDEX MODELS

1 SUPPOSE $F \subset \mathbb{R}^S$! $R \text{ AND } A_2 \in \mathbb{R}^D$ $I = \dots; S$

$$F(X) = G(h_{A_1}; X; \dots; h_{A_S}; X); \quad X \in \mathbb{R}^D \text{ (; 2):}$$

1 $\mathcal{S} := \text{SPAN}(A_1; \dots; A_S)$ \mathbb{R}^D IS THE S DIMENSIONAL RELEVANT FEATURE SUBSPACE

1 \hat{F}_N BE A (S) 4 FOREST ESTIMATOR WITH DIRECTIONAL DISTRIBUTION

$$N = (\dots)_3 + (\dots)_{N^3};$$

FOR $N \in (;)$ WHERE $\text{SUPP}(P_1) = \mathcal{S} \setminus \mathcal{S}^D$ $\text{SUPP}(P_3) = \mathcal{S}^? \setminus \mathcal{S}^D$

2 RATES FOR MULTI INDEX MODELS

1) SUPPOSE \mathbb{R}^S ! $R \text{ AND } A_2 \text{ R } D \text{ I} = ; \dots ; S$

$$f(x) = G(h_{A_1}; x; \dots; h_{A_S}; x); \quad x \in \mathbb{R}^D \text{ (; 2):}$$

2) $\mathbb{R}^3 := \text{SPAN}(A_1; \dots; A_S)$ \mathbb{R}^D IS THE S DIMENSIONAL RELEVANT FEATURE SUBSPACE

3) $\hat{F}_{N, N-}$ BE A $(3, 4)$ FOREST ESTIMATOR WITH DIRECTIONAL DISTRIBUTION

$$N = (\quad "N)_3 + "N_3^?$$

FOR $N \in (;)$ WHERE $\text{SUPP}(P) = \mathbb{R}^3 \setminus S^D$ $\text{SUPP}(P_3) = \mathbb{R}^3 \setminus S^D$

4 THEOREM

1) $(F_{N, N-})$ IS IPSCHITZ LETTING $N = (S^?)$ AND $N = (S^?)$ GIVES

$$E[(\hat{F}_{N, N-}(8) - f(8))] / N = (S^?)$$

2) ! SIMILAR EXTENSION HOLDS IF

PROOF OF THE VARIANCE DECOMPOSITION OF THE

WE HAVE THE FOLLOWING VARIANCE DECOMPOSITION

$$E[(\hat{F}_N(t) - F(t))^2] = E[(F(t) - \bar{F}(t))^2] + E[(\bar{F}(t) - \hat{F}_N(t))^2];$$

WHERE

$$\bar{F}(x) = E_8[F(t) | x]; \quad x \in \mathcal{X};$$

IS THE ORTHOGONAL PROJECTION OF F ONTO THE SUBSPACE OF FUNCTIONS THAT ARE CONSTANT WITHIN THE CELL OF THE 3^4 TESSELLATION

PROOF OF THE VARIANCE DECOMPOSITION OF THE

- WE HAVE THE FOLLOWING VARIANCE DECOMPOSITION

$$E[(\hat{F}_N(x) - F(x))^2] = E[(F(x) - F(x))^2] + E[(F(x) - \hat{F}_N(x))^2];$$

WHERE

$$F(x) = E[F(x) | X \in C]; \quad X \in C;$$

IS THE ORTHOGONAL PROJECTION OF $F(x)$ ONTO THE SUBSPACE OF FUNCTIONS THAT ARE CONSTANT WITHIN THE CELL OF THE h -TESSELLATION

- THIS IS CONTROLLED BY THE SIZE OF THE CELL CONTAINING x
- VARIANCE IS CONTROLLED BY THE EXPECTED NUMBER OF CELLS

3 STATIONARY 34)4 4 ESSELLATION ON

- | $Y(\cdot; \gamma) \subset \mathbb{R}^D$ IN COMPACT AND CONVEX
 $\gamma \in \mathbb{R}^D$ WITH LIFETIME
- | CONSISTENCY FOR γ

$$Y(\cdot; \gamma) \stackrel{D}{=} Y(\cdot; \gamma) \setminus \gamma$$

STATIONARY 34)4 TESSELLATION ON \mathbb{R}^D

- | $Y(\cdot; \gamma) \subset \mathbb{R}^D$ IN COMPACT AND CONVEX
 $\gamma \subset \mathbb{R}^D$ WITH LIFETIME
- | CONSISTENCY FOR γ

$$Y(\cdot; \gamma) \stackrel{D}{=} Y(\cdot; \gamma) \setminus \gamma$$

- | THERE EXISTS STATIONARY TESSELLATION ON \mathbb{R}^D SUCH THAT
 - | $Y(\cdot) \setminus \gamma \stackrel{D}{=} Y(\cdot; \gamma)$ FOR ALL COMPACT

3 STATIONARY 34)4 TESSELLATION ON \mathbb{R}^D

- | $Y(\cdot; \gamma) \subset \mathbb{R}^D$ IN COMPACT AND CONVEX SETS WITH LIFETIME

- | CONSISTENCY FOR γ

$$Y(\cdot; \gamma) \stackrel{D}{=} Y(\cdot; \gamma) \setminus \gamma$$

- | THERE EXISTS STATIONARY TESSELLATION ON \mathbb{R}^D SUCH THAT

- | $Y(\cdot) \setminus \gamma \stackrel{D}{=} Y(\cdot; \gamma)$ FOR ALL COMPACT

- | TABLE UNDER ITERATION

$$Y(\cdot) \stackrel{D}{=} N(Y(\cdot), Y(\cdot));$$

$$\text{STY } Y := Y \left[\overset{S}{\cup} \text{ } \right]_{\text{CELLS}} (Y(Q) \setminus Q)$$

WHERE $Y(Q) : \mathbb{C}^2$ ARE IID COPIES OF

- | SCALING PROPERTY $Y(\cdot) \stackrel{D}{=} Y(\cdot)$

#ELLS OF STATIONARY RANDOM TESSELLATIONS

- | #ELLS OF \mathcal{K}^d) FORM A STATIONARY POINT PROCESS ON SPACE OF COMPACT CONVEX POLYTOPES

#CELLS OF STATIONARY RANDOM TESSELLATIONS

1 #CELLS OF (\mathcal{X}) FORM A STATIONARY POINT PROCESS ON SPACE COMPACT CONVEX POLYTOPES

2 TYPICAL CELL CENTERED RANDOM POLYTOPES SUCH THAT FOR ALL (K)

$$E \sum_{x \in \mathcal{X}} \int_{K(x)} \text{vol}_D(\cdot) = \frac{E \int_{\mathbb{R}^D} \text{vol}_D(\cdot)}{E \int_{\mathbb{R}^D} \text{vol}_D(\cdot)}$$

PROOF) DECENTRALIZED VARIANCE BOUND

- DECENTRALIZED VARIANCE IS CONTROLLED BY THE EXPECTED NUMBER OF CELLS
- LET \mathcal{D} BE A COMPACT AND CONVEX SET
- LET $\mathcal{Y}(\cdot)$ BE A 3D TESSERATION WITH FINITE LIFETIME

EMMA

LET $\mathcal{Y}(\cdot)$ BE THE TYPICAL CELL THEN

$$E[\text{vol}_D(\mathcal{Y}(\cdot))] = \frac{E[\sum_{K \in \mathcal{Y}(\cdot)} \text{vol}_D(K)]}{E[\#\text{cells}(\mathcal{Y}(\cdot))]}$$

WHERE $E[\sum_{K \in \mathcal{Y}(\cdot)} \text{vol}_D(K)] := E[\sum_{K \in \mathcal{Y}(\cdot)} \sum_{D \in \mathcal{K}} \text{vol}_D(K)]$

PROOF) DECENT VARIANCE FOUND

- DECENT VARIANCE IS CONTROLLED BY THE EXPECTED NUMBER OF CELLS
- LET R^D BE A COMPACT AND CONVEX SET
- LET (Y_i) BE A 3D TESSELLATION FROM WITH LIFETIME

EMMA

LET μ BE THE TYPICAL CELL VOLUME

$$E[\sum_{i \in \mathcal{Z}^D(Y_i)} \mu] = \sum_{K \in \mathcal{K}} \frac{E[\mu(K; [D, K])]}{E[\text{vol}_D(\cdot)]}$$

WHERE $\mu(K; [D, K]) := E[\mu(\{z\}_K; \{z\}_{D-K})]$

- PROOF APPLY STEINER'S FORMULA TO $\sum_{i \in \mathcal{Z}^D(Y_i)} \mu = \frac{E[\text{vol}_D(\cdot)]}{E[\text{vol}_D(\cdot)]}$

PROOF OF DEA 2ISK FOUND

- Let \hat{F}_N be the forest estimator corresponding to the (34)
- Let \mathcal{C} be the zero cell containing the origin and $\mathcal{C}(L)$ of

ASSUME: γ is Lipschitz then

$$E[\hat{F}_N(\gamma) - F(\gamma)]$$

$$\leq \frac{E[\text{diam}(\mathcal{C})]}{N} + \frac{(k_1 + k_2) X^D}{N} \leq \frac{D}{K} \frac{E[\gamma(\mathcal{C})]}{E[\text{vol}_D(\mathcal{C})]}$$

- Setting $N = N^{(D+1)}$ gives the minimax rate for Lipschitz function

PROOF OF THE BOUND FOR THE INDEX OF THE

1. SUPPOSE $f \in \mathcal{C}^s$! RANDOM $X_1, \dots, X_n \in \mathbb{R}^D$

$$f(x) := G(h_{A_1}; x, \dots, h_{A_s}; x)$$

2. $\mathcal{S} := \text{SPAN}(A_1, \dots, A_s)$ \mathbb{R}^D IS THE s -DIMENSIONAL RELEVANT FEATURE SUBSPACE

3. \hat{f}_n BE A s -D FOREST ESTIMATOR WITH DIRECTIONAL DISTRIBUTION

$$N = (\quad)^s + N^s?$$

PROOF OF THE BOUND FOR THE INDEX OF THE

1. SUPPOSE \mathcal{F} IS A RANDOM SUBSET OF S

$$f(x) := G(h_A; x; \dots; h_{A_S}; x)$$

2. $S := \text{SPAN}(A_1; \dots; A_S)$ IS THE D -DIMENSIONAL RELEVANT FEATURE SUBSPACE

3. \mathcal{F} IS A $(1/\epsilon)$ -FOREST ESTIMATOR WITH DIRECTIONAL DISTRIBUTION

$$N = (\dots) \epsilon + N \epsilon^2$$

4. ASSUME \mathcal{F} IS $(1/\epsilon)$ -IPSCHITZ THEN

$$E[\| \hat{F}_N(\delta) - F(\delta) \|] \leq \frac{C \cdot k \cdot k_{OP}}{N} E[\text{diam}(\mathcal{F} \setminus S)] + \frac{(C k_1 + \dots)}{N} \left(\frac{D}{N} \right)^{2D} \left(\frac{1}{\epsilon} \right)^{D-1} + C \left(\frac{D}{N} \right)^{D-1} :$$

5. S AND S^c ARE CONVEX BODIES ZONED BY S

SUMMARY AND FUTURE WORK

- | WE HAVE PROVED MINIMAX OPTIMAL RATES FOR A LARGE CLASS OF RANDOM FOREST PARTITION ESTIMATORS WITH GENERAL SPLIT DIRECTIONS
- | THEORY OF STATIONARY RANDOM TESSELLATIONS AND A FRAMEWORK FOR UNDERSTANDING AND DEVELOPING RANDOM PARTITIONS
- | PERFORMANCE DEPENDS ON GEOMETRY OF THE CELLS AND VOLUMES OF TYPICAL CELL

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- | HOW TO LEARN DIRECTIONAL DISTRIBUTION FROM DATA
- | OTHER APPLICATIONS CLUSTERING RANDOM FEATURE MODELS

Application #2: Optimal regularizers for a data source

Joint work with Oscar Leong (Caltech), Yong Sheng Soh (National University of Singapore), and Venkat Chandrasekaran (Caltech)

Inverse Problems and regularization

- | Goal is to recover signal x from:

$$y = A(x) + \epsilon$$

where A is a known forward map and ϵ is observation noise

- | Problem may be *ill-posed*

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- | Problem may be *ill-posed*

Functional Analytic Regularization:

- | Recover x with the following optimization problem:

$$\operatorname{argmin}_x \operatorname{loss}(A(x); y) + \text{regularizer}(x)$$

- | Regularizer function promotes structure in the solution

Variety of regularizers

Hand-crafted:

Sparsity is promoted by ℓ_1 norm (convex) and by ℓ_p norm for $p \geq 1$ (non-convex)

$$\operatorname{argmin}_x \operatorname{loss}(A(x); y) + \|x\|_p$$

Variety of regularizers

Hand-crafted:

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Data-driven:

Dictionary Learning (*convex*)

- Learn $A \in \mathbb{R}^{d \times p}$ such that $x = Az$, where z is sparse

$$\|A^T x\|_1 \quad (\cdot) \quad \|x\|_{A(B_1)}$$

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Generative models (**non-convex**)

- Neural network based regularization

Main question considered in this work

Which regularizer should one choose?

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Which regularizer should one choose?

- | What is the optimal regularizer to impose for a given data source?
- | Convex versus nonconvex?

Set-up and assumptions

- | Let P be a probability distribution on \mathbb{R}^d modeling a data source
- | Define optimal regularizer f from a family F as a solution to:

$$\operatorname{argmin}_{f \in F} E_P[f(x)]$$

- | Conditions on $f \in F$:
 - | Positively homogenous: $f(\alpha x) = \alpha f(x), \quad \alpha \geq 0$
 - | $f \geq 0$ and continuous

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$f \in F \iff f = k \cdot k_K$ is the Minkowski functional of a **star body** K

Minkowski functional of a compact set $C \subset \mathbb{R}^d$:

$$k_C(x) := \inf \{t > 0 : x \in tC\}$$

Star bodies and radial functions

- | The **radial function** of a compact set $K \subset \mathbb{R}^d$ is defined by

$$\kappa(x) := \sup \{ t > 0 : tx \in K \} = \|x\|_K^{-1}$$

- | A compact set $K \subset \mathbb{R}^d$ is a **star body** if κ is continuous and it is *starshaped* (with respect to the origin)

$$x \in K \implies [0; x] \subset K$$

- | Star bodies are uniquely determined by their radial functions

Unique optimal regularizer

Theorem (Leong, Soh, Chandrasekaran, O., 2022+)

Let P be a distribution on \mathbb{R}^d with density p and assume $E_P[\|x\|_2] < 1$.

Unique optimal regularizer

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Let P be a distribution on \mathbb{R}^d with density p and assume $E_P[\|x\|_2] < 1$. Suppose the following function is continuous:

$$p(u) := \int_0^1 r^{d-1} p(ru) dr^{d-1}; \quad u \in S^{d-1}. \quad (1)$$

Unique optimal regularizer

Theorem (Leong, Soh, Chandrasekaran, O., 2022+)

Let P be a distribution on \mathbb{R}^d with density p and assume $E_P[\|x\|_2] < \infty$. Suppose the following function is continuous:

$$\rho(u) := \int_0^{\|u\|_2^{-1}} r^d p(ru) dr^{1-(d+1)}; \quad u \in S^{d-1}. \quad (1)$$

Then there is a unique star body L_P with radial function ρ , and

$$K := \text{vol}_d(L_P)^{1/d} L_P$$

is the unique solution to

$$\operatorname{argmin}_{K \in S^d: \text{vol}_d(K)=1} E_P[\|x\|_K]$$

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is the unique solution to

$$\underset{K \in S^d: \text{vol}_d(K)=1}{\text{argmin}} E_P[\|x\|_K]$$

- 1 If L_P is convex, then the optimal regularizer is convex!

Examples

(i) Densities induced by star bodies:

$$p(x) = (kxk_L) \quad L_p = c \quad L$$

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(i) Densities induced by star bodies:

$$p(x) = \frac{1}{L_P} \int_0^{\|x\|} \phi\left(\frac{x}{\|x\|} t\right) dt$$

(ii) Gaussian Mixtures:

$$P = \frac{1}{2} N(0; \mu_1) + \frac{1}{2} N(0; \mu_2)$$

where $\mu_1 := [1; 0; 0; \dots] \in \mathbb{R}^2$ and $\mu_2 := [0; 0; 0; 1] \in \mathbb{R}^2$

Figure: Plots of L_P for $\epsilon = 0.1$ (left) and $\epsilon = 0.01$ (right).

Proof

Goal: Characterize unique solution to

$$\operatorname{argmin}_{K \subseteq S^d: \operatorname{vol}_d(K)=1} E_P[\|x\|_K]$$

By change to polar coordinates,

$$\begin{aligned} E_P[\|x\|_K] &= \int_{\mathbb{R}^d} \|x\|_K p(x) dx = \int_{S^{d-1}} \int_0^{\infty} \|u\|_K r^d p(ru) dr du \\ &= \int_{S^{d-1}} \|u\|_K^{-1} \rho(u)^{d+1} du := d \int_{S^{d-1}} \rho(u)^{d+1} du \end{aligned}$$

Proof

Goal: Characterize unique solution to

$$\operatorname{argmin}_{K \subset S^d: \operatorname{vol}_d(K)=1} E_P[\|x\|_K]$$

By change to polar coordinates,

$$\begin{aligned} E_P[\|x\|_K] &= \int_{S^{d-1}} \int_0^{\rho_K(u)} \int_{S^{d-1}} \|x\|_K p(x) dx = \int_{S^{d-1}} \int_0^{\rho_K(u)} r^d p(ru) dr du \\ &= \int_{S^{d-1}} \rho_K(u)^{d+1} p(u) du := dV_{-1}(K; L_P) \end{aligned}$$

Theorem (Dual Mixed Volume Inequality (Lutwak, 1975))

For star bodies K and L ,

$$V_{-1}(K; L)^d \leq \operatorname{vol}_d(K)^{d-1} \operatorname{vol}_d(L)^{d+1};$$

and equality holds if and only if L and K are dilates, i.e. $L = \lambda K$ for some $\lambda > 0$

Summary and future work

- | Dual Brunn-Minkowski theory provides tools for characterizing optimal functional for imposing structure on a dataset for inverse problems
- | Other results: convergence of empirical minimizers and generalization error bounds

Summary and future work

- | Dual Brunn-Minkowski theory provides tools for characterizing optimal functional for imposing structure on a dataset for inverse problems
- | Other results: convergence of empirical minimizers and generalization error bounds

- | How do optimal regularizers perform in downstream tasks?
- | How to efficiently compute the optimal regularizer?

Papers

“Minimax Rates for High-Dimensional Random Tessellation Forests”
Joint with Ngoc Mai Tran. <https://arxiv.org/abs/2109.10541>

“Optimal Convex and Nonconvex Regularizers for a Data Source”
Joint with Oscar Leong, Yong Sheng Soh, and Venkat Chandrasekaran. In preparation.

Thank you!